

# Heavy-Tailed Longitudinal Linear Mixed Models for Multiple Censored Responses Data

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## Motivation

- ▶ In AIDS studies it is quite common to observe viral load measurements that are collected irregularly over time. Moreover, these measurements can be subjected to some upper and/or lower detection limits depending on the quantification assays.
- ▶ A complication arises when these continuous repeated measures have a heavy-tailed behavior.
- ▶ For such data structures, we propose a robust censored linear mixed model for multiple responses based on the class of multivariate scale mixtures of normal distributions.

# Motivating data

A5055 clinical trial

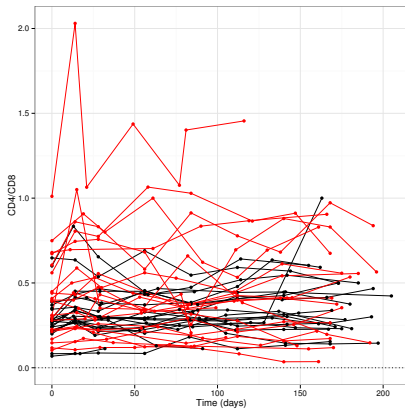
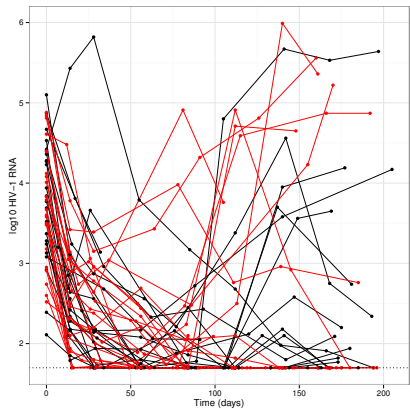
## ► A5055 study

- 44 infected patients with the human immunodeficiency virus type 1 (HIV-1) treated with one of the two potent antiretroviral (ARV) therapies;
- 2 outcomes:  $\log_{10}(\text{RNA})$  and CD4/CD8, where CD4 and CD8 are two immunologic markers frequently used for monitoring disease progression in AIDS studies;
- 33% (106 out of 316) of measurements lying below the limits of assay quantification (left-censored).

# Motivating data

A5055 clinical trial

◀ A5055



## Recent works

### Longitudinal Models

Censored longitudinal models with heavy-tailed distribution

- ▶ Lachos *et al.* (2011) [*Biometrics*]
- ▶ **Garay et al. (2014)** [*Statistical Methods in Medical Research*]
- ▶ **Matos et al. (2013b)** [*Statistica Sinica*]
- ▶ **Wang et al. (2015)** [*Statistical Methods in Medical Research*]

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# Scale mixture of normal distributions (SMN)

Andrews & Mallows (1974); Lange & Sinsheimer (1993)

## Stochastic representation

$$\mathbf{y} = \boldsymbol{\mu} + \kappa(U)^{1/2}\mathbf{Z}, \quad (1)$$

where,

- ▶  $\boldsymbol{\mu} \in \mathbb{R}$  is a location vector;
- ▶  $Z \sim N(0, \boldsymbol{\Sigma})$ ;
- ▶  $U$  is a positive random variable with cumulative distribution function (cdf)  $H(u|\boldsymbol{\nu})$  and probability density function (pdf)  $h(u|\boldsymbol{\nu})$  independent of  $Z$ ;
- ▶  $\kappa(U)$  is the weight function;
- ▶ Notation:  $\mathbf{y} \sim \text{SMN}_p(\boldsymbol{\mu}, \boldsymbol{\Sigma}; H)$ .

$$\begin{aligned} \mathbf{y}|U = u &\sim N(\boldsymbol{\mu}, \kappa(u)\boldsymbol{\Sigma}), \\ U = u &\sim h(u|\boldsymbol{\nu}). \end{aligned} \quad (2)$$

# Scale mixture of normal distributions (SMN)

Special cases:  $\mathbf{y} \in \mathbb{R}^p$  and  $\kappa(u) = 1/u$ ;

- ▶ **The multivariate normal distribution**

- ▶ Distribution function:  $N(\mathbf{y}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \phi_p(\mathbf{y}; \boldsymbol{\mu}, \boldsymbol{\Sigma})$ .

- ▶ **The multivariate Student's-t distribution**

- ▶ Distribution function:

$$T(\mathbf{y}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \nu) = \frac{\Gamma(\frac{p+\nu}{2})}{\Gamma(\frac{\nu}{2})\pi^{p/2}} \nu^{-p/2} |\boldsymbol{\Sigma}|^{-1/2} \left(1 + \frac{d}{\nu}\right)^{-(p+\nu)/2}.$$

- ▶ **The multivariate slash distribution**

- ▶ Distribution function:

$$SL(\mathbf{y}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \nu) = \nu \int_0^1 u^{\nu-1} \phi_p(\mathbf{y}; \boldsymbol{\mu}, u^{-1}\boldsymbol{\Sigma}) du, \quad u \in (0, 1), \quad \nu > 0.$$

- ▶ **The multivariate contaminated normal distribution**

- ▶ Distribution function:

$$CN(\mathbf{y}|\boldsymbol{\mu}, \boldsymbol{\Sigma}, \nu) = \nu \phi_p(\mathbf{y}; \boldsymbol{\mu}, \gamma^{-1}\boldsymbol{\Sigma}) + (1 - \nu) \phi_p(\mathbf{y}; \boldsymbol{\mu}, \boldsymbol{\Sigma}).$$

# SAEM Algorithm

SAEM Algorithm - Delyon *et al.* (1999)

Let  $\theta$  be the parameter vector and  $\mathbf{y}_c = (\mathbf{y}^\top, \mathbf{q}^\top)$  be the vector of complete data, i.e., the observed data  $\mathbf{y}^\top$  and the missing/censored data (or the latent variables, depending on the situation)  $\mathbf{q}^\top$ . The SAEM algorithm consists in:

► **E-Step:**

1. **Simulation-step:** Draw  $\mathbf{q}^{(k,l)}$  ( $l = 1, \dots, m$ ) from the conditional distribution  $f(\mathbf{q}|\mathbf{y}, \hat{\theta}^{(k-1)})$ ;

2. **Stochastic-approximation-step:** Update  $Q(\theta|\hat{\theta}^{(k)})$  according to

$$Q(\theta|\hat{\theta}^{(k)}) = Q(\theta|\hat{\theta}^{(k-1)}) + \delta_k \left[ \frac{1}{m} \sum_{l=1}^m \ell_c(\theta|\mathbf{q}^{(k,l)}, \mathbf{y}) - Q(\theta|\hat{\theta}^{(k-1)}) \right],$$

where  $\ell_c(\theta | \mathbf{y}_c) = \sum_{i=1}^n \ell_i(\theta | \mathbf{y}_c)$  is the complete log-likelihood function and  $\delta_k$  is a smoothness parameter, i.e., a decreasing sequence of positive numbers such that  $\sum_{k=1}^{\infty} \delta_k = \infty$  and  $\sum_{k=1}^{\infty} \delta_k^2 < \infty$ .

# SAEM Algorithm

SAEM Algorithm - Delyon *et al.* (1999)

- ▶ **M-Step:** Update  $\theta^{(k)}$  according to

$$\hat{\theta}^{(k+1)} = \underset{\theta}{\operatorname{argmax}} Q(\theta | \hat{\theta}^{(k)}).$$

- ▶ As proposed by Galarza *et al.* (2015), we will consider the following smoothing parameter

$$\delta_k = \begin{cases} 1, & \text{if } 1 \leq k \leq cW; \\ \frac{1}{k-cW}, & \text{if } cW + 1 \leq k \leq W, \end{cases} \quad (3)$$

where,

- $W$  is the maximum number of iterations; e
- $c$  is a cut point ( $0 \leq c \leq 1$ ) which determines the percentage of the initial iterations.

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## SMN-MLMEC model

Let  $\mathbf{Y}_i = [\mathbf{y}_{i1} : \dots : \mathbf{y}_{ir}]$ , then

$$\mathbf{y}_i = \mathbf{X}_i \boldsymbol{\beta} + \mathbf{Z}_i \mathbf{b}_i + \boldsymbol{\epsilon}_i, \quad i = 1, \dots, n; \quad (4)$$

where:

- ▶  $\mathbf{y}_i = \text{vec}(\mathbf{Y}_i) = (\mathbf{y}_{i1}^\top, \dots, \mathbf{y}_{ir}^\top)^\top$ , where  $\mathbf{y}_{ij} = (y_{ij1}, \dots, y_{ijn_i})^\top$  is a  $n_i \times 1$  vector of the  $j$ th outcome for the  $i$ th subject;
- ▶  $\mathbf{X}_i = \text{Bdiag}\{\mathbf{X}_{i1}, \dots, \mathbf{X}_{ir}\}$ , where  $\mathbf{X}_{ij}$  is an  $n_i \times p_j$  design matrix for fixed effects corresponding to the  $j$ th outcome of the  $i$ th subject;
- ▶  $\mathbf{Z}_i = \text{Bdiag}\{\mathbf{Z}_{i1}, \dots, \mathbf{Z}_{ir}\}$ , onde  $\mathbf{Z}_{ij}$  is an  $n_i \times q_j$  design matrix for random effects corresponding to the  $j$ th outcome of the  $i$ th subject, which is generally a subset of  $\mathbf{X}_{ij}$ ;
- ▶  $\boldsymbol{\beta} = (\boldsymbol{\beta}_1^\top, \dots, \boldsymbol{\beta}_r^\top)^\top$  is the  $p \times 1$  vector of fixed effects associated with the design matrix  $\mathbf{X}_i$ ,  $p = \sum_{j=1}^r p_j$ ;
- ▶  $\mathbf{b}_i = (\mathbf{b}_{i1}^\top, \dots, \mathbf{b}_{ir}^\top)^\top$  is the  $q \times 1$  vector of random effects associated with the design matrix  $\mathbf{Z}_i$ ,  $q = \sum_{j=1}^r q_j$ ;
- ▶  $\boldsymbol{\epsilon}_i = \text{vec}(\mathbf{E}_i) = (\boldsymbol{\epsilon}_{i1}^\top, \dots, \boldsymbol{\epsilon}_{ir}^\top)^\top$  is the vector of random errors of dimension  $(s_i \times 1)$  ( $s_i = n_i \times r$ ), where  $\mathbf{E}_i = [\boldsymbol{\epsilon}_{i1} : \dots : \boldsymbol{\epsilon}_{ir}]$  and  $\boldsymbol{\epsilon}_{ij}$  corresponds to the error for the  $j$ th outcome for the  $i$ th subject.

## SMN-MLMEC model

- ▶ Considering the multivariate SMN distributions for the random terms, the model can be expressed as

$$\begin{aligned} \mathbf{y}_i \mid \mathbf{b}_i &\stackrel{\text{ind.}}{\sim} \text{SMN}_{s_i}(\mathbf{X}_i\boldsymbol{\beta} + \mathbf{Z}_i\mathbf{b}_i, \mathbf{R}_i; H_1), \\ \mathbf{b}_i &\stackrel{\text{ind.}}{\sim} \text{SMN}_q(\mathbf{0}, \mathbf{D}; H_2), \quad i = 1, \dots, n. \end{aligned} \quad (5)$$

- ▶ Using the stochastic representation (1), the hierarchical representation of the model defined in(4) is given by

$$\begin{aligned} \mathbf{y}_i \mid \mathbf{b}_i, \kappa_i &\stackrel{\text{ind.}}{\sim} N_{s_i}(\mathbf{X}_i\boldsymbol{\beta} + \mathbf{Z}_i\mathbf{b}_i, \kappa_i^{-1}\mathbf{R}_i), \\ \mathbf{b}_i \mid \tau_i &\stackrel{\text{ind.}}{\sim} N_q(\mathbf{0}, \tau_i^{-1}\mathbf{D}), \\ \kappa_i &\stackrel{\text{ind.}}{\sim} H_1(\nu), \\ \tau_i &\stackrel{\text{ind.}}{\sim} H_2(\eta); \end{aligned} \quad (6)$$

where  $\mathbf{R}_i = \boldsymbol{\Sigma} \otimes \boldsymbol{\Omega}_i$ .

## SMN-MLMEC model

Correlation structures - Muñoz *et al.* (1992)

Damped exponential correlation (DEC):

$$\Omega_i = \Omega_i(\phi, \mathbf{t}_i) = \left[ \phi_1^{|\mathbf{t}_{ij} - \mathbf{t}_{ik}|^{\phi_2}} \right], \quad i = 1, \dots, n, \quad j, k = 1, \dots, n_i, \quad (7)$$

For the DEC structure, we have that:

- (a) if  $\phi_2 = 0$ , then  $\Omega_i$  generates the compound symmetry correlation structure;
- (b) when  $0 < \phi_2 < 1$ , then  $\Omega_i$  presents a decay rate between the compound symmetry structure and the first-order AR (AR (1)) model;
- (c) if  $\phi_2 = 1$ , then  $\Omega_i$  generates an AR(1) structure;
- (d) when  $\phi_2 > 1$ ,  $\Omega_i$  presents a decay rate faster than the AR(1) structure; and
- (e) if  $\phi_2 \rightarrow \infty$ , then  $\Omega_i$  represents the first-order moving average model, MA(1).



## SMN-MLMEC model

- ▶ Recall that we are interested in the case where left-censored observations can occur. That is, we assume that the observations are of the form

$$\begin{aligned} y_{ijk} &\leq V_{ijk} \quad \text{se } C_{ijk} = 1, \\ y_{ijk} &= V_{ijk} \quad \text{se } C_{ijk} = 0, \end{aligned}$$

with  $i = 1, \dots, n$ ,  $j = 1, \dots, n_i$  and  $k = 1, \dots, r$ ;

- ▶ The observed data for the  $i$ -th subject is represented by  $(\mathbf{V}_i, \mathbf{C}_i)$ , where  $\mathbf{V}_i = [V_{i1} : \dots : V_{ir}]$  is an  $n_i \times r$  matrix and  $\mathbf{C}_i = [C_{i1} : \dots : C_{ir}]$  is an  $n_i \times r$  matrix;
- ▶ The extensions to arbitrary censoring are immediate.

# SMN-MLMEC model

## Maximum likelihood estimation - SAEM

- ▶ The complete data log-likelihood function:
  - ▶  $\theta = (\beta, \sigma, \alpha, \phi, \nu, \eta)$ ;
  - ▶ Augmenting data:  $\mathbf{y}_c = (\mathbf{V}^\top, \mathbf{C}^\top, \mathbf{y}^\top, \mathbf{b}^\top, \boldsymbol{\kappa}^\top, \boldsymbol{\tau}^\top)^\top$ ;
  - ▶  $[\mathbf{V} \ \mathbf{C} \ \mathbf{y}] \Rightarrow [\mathbf{y}]$ .

$$\begin{aligned}\ell_c(\theta|\mathbf{y}_c) &= \sum_{i=1}^n [\log f(\mathbf{y}_i|\mathbf{b}_i, \kappa_i) + \log f(\mathbf{b}_i|\tau_i) + \log h_1(\kappa_i|\nu) + \log h_2(\tau_i|\eta)] \\ &= -\frac{1}{2} \sum_{i=1}^n \log |\mathbf{R}_i| - \frac{1}{2} \sum_{i=1}^n \kappa_i (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta} - \mathbf{Z}_i \mathbf{b}_i)^\top \mathbf{R}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta} - \mathbf{Z}_i \mathbf{b}_i) \\ &\quad - \frac{1}{2} \sum_{i=1}^n \log |\mathbf{D}| - \frac{1}{2} \sum_{i=1}^n \tau_i \mathbf{b}_i^\top \mathbf{D}^{-1} \mathbf{b}_i + \sum_{i=1}^n \log h_1(\kappa_i|\nu) + \sum_{i=1}^n \log h_2(\tau_i|\eta) + K,\end{aligned}$$

with  $K$  being a constant that does not depend on the parameter vector  $\theta$ .

# SMN-MLMEC model

## Maximum likelihood estimation - SAEM

- Q-function: For the  $i$ -th subject,

$$\begin{aligned}
 Q_i(\boldsymbol{\theta} | \widehat{\boldsymbol{\theta}}^{(k)}) &= \widehat{\ell h_{1i}}^{(k)} + \widehat{\ell h_{2i}}^{(k)} - \frac{1}{2} \log |\widehat{\mathbf{D}}^{(k)}| - \frac{1}{2} \text{tr} \left( \widehat{\boldsymbol{\tau}} \widehat{\mathbf{b}}_i^{2(k)} \widehat{\mathbf{D}}_i^{-1(k)} \right) - \frac{1}{2} \sum_{i=1}^n \log |\widehat{\mathbf{R}}_i^{(k)}| \\
 &- \frac{1}{2} \left[ \text{tr} \left( \widehat{\boldsymbol{\kappa}} \widehat{\mathbf{y}}_i^{2(k)} \widehat{\mathbf{R}}_i^{-1(k)} \right) - 2 \widehat{\boldsymbol{\beta}}^{(k)\top} \mathbf{X}_i^\top \widehat{\mathbf{R}}_i^{-1(k)} \widehat{\boldsymbol{\kappa}} \widehat{\mathbf{y}}_i^{(k)} + 2 \widehat{\boldsymbol{\beta}}^{(k)\top} \mathbf{X}_i^\top \widehat{\mathbf{R}}_i^{-1(k)} \mathbf{Z}_i \widehat{\boldsymbol{\kappa}} \widehat{\mathbf{b}}_i^{(k)} \right. \\
 &- \left. 2 \text{tr} \left( \mathbf{Z}_i^\top \widehat{\mathbf{R}}_i^{-1(k)} \widehat{\boldsymbol{\kappa}} \widehat{\mathbf{y}}_i \widehat{\mathbf{b}}_i^{(k)} \right) + \text{tr} \left( \mathbf{Z}_i^\top \widehat{\mathbf{R}}_i^{-1(k)} \mathbf{Z}_i \widehat{\boldsymbol{\kappa}} \widehat{\mathbf{b}}_i^{2(k)} \right) + \widehat{\kappa}_i^{(k)} \widehat{\boldsymbol{\beta}}^{(k)\top} \mathbf{X}_i^\top \widehat{\mathbf{R}}_i^{-1(k)} \mathbf{X}_i \widehat{\boldsymbol{\beta}}^{(k)} \right],
 \end{aligned}$$

with

$$\begin{aligned}
 \widehat{\ell h_{1i}}^{(k)} &= E \left[ \log h_1(\kappa_i | \nu) | \mathbf{V}_i, \mathbf{C}_i, \widehat{\boldsymbol{\theta}}^{(k)} \right], & \widehat{\ell h_{2i}}^{(k)} &= E \left[ \log h_2(\tau_i | \eta) | \mathbf{V}_i, \mathbf{C}_i, \widehat{\boldsymbol{\theta}}^{(k)} \right] \\
 \widehat{\boldsymbol{\kappa}} \widehat{\mathbf{y}}_i^{2(k)} &= E \left[ \boldsymbol{\kappa}_i \mathbf{y}_i \mathbf{y}_i^\top | \mathbf{V}_i, \mathbf{C}_i, \widehat{\boldsymbol{\theta}}^{(k)} \right], & \widehat{\boldsymbol{\kappa}} \widehat{\mathbf{y}}_i^{(k)} &= E \left[ \boldsymbol{\kappa}_i \mathbf{y}_i | \mathbf{V}_i, \mathbf{C}_i, \widehat{\boldsymbol{\theta}}^{(k)} \right], \\
 \widehat{\boldsymbol{\kappa}} \widehat{\mathbf{b}}_i^{2(k)} &= E \left[ \boldsymbol{\kappa}_i \mathbf{b}_i \mathbf{b}_i^\top | \mathbf{V}_i, \mathbf{C}_i, \widehat{\boldsymbol{\theta}}^{(k)} \right], & \widehat{\boldsymbol{\kappa}} \widehat{\mathbf{b}}_i^{(k)} &= E \left[ \boldsymbol{\kappa}_i \mathbf{b}_i | \mathbf{V}_i, \mathbf{C}_i, \widehat{\boldsymbol{\theta}}^{(k)} \right], \\
 \widehat{\boldsymbol{\tau}} \widehat{\mathbf{b}}_i^{2(k)} &= E \left[ \boldsymbol{\tau}_i \mathbf{b}_i \mathbf{b}_i^\top | \mathbf{V}_i, \mathbf{C}_i, \widehat{\boldsymbol{\theta}}^{(k)} \right], & \widehat{\boldsymbol{\kappa}} \widehat{\mathbf{y}}_i \widehat{\mathbf{b}}_i^{(k)} &= E \left[ \boldsymbol{\kappa}_i \mathbf{y}_i \mathbf{b}_i^\top | \mathbf{V}_i, \mathbf{C}_i, \widehat{\boldsymbol{\theta}}^{(k)} \right], \\
 \widehat{\kappa}_i^{(k)} &= E \left[ \boldsymbol{\kappa}_i | \mathbf{V}_i, \mathbf{C}_i, \widehat{\boldsymbol{\theta}}^{(k)} \right].
 \end{aligned}$$

# SMN-MLMEC model

## SAEM - E-step

► **Simulation-step:** Gibbs Sampler

**Step 1.** Sample  $\mathbf{y}_i^c$  de  $f(\mathbf{y}_i^c | \mathbf{V}_i^c, \mathbf{y}_i^o, \mathbf{b}_i^{(k,l-1)}, \kappa_i^{(k,l-1)}, \tau_i^{(k,l-1)}, \hat{\boldsymbol{\theta}}^{(k)})$ , where

$$\mathbf{y}_i^c | \mathbf{V}_i^c, \mathbf{y}_i^o, \mathbf{b}_i, \kappa_i, \tau_i, \boldsymbol{\theta} \sim \text{TN}_{s_i^c}(\boldsymbol{\mu}_i, \kappa_i^{-1} \mathbf{S}_i; \mathbb{A}_i),$$

with  $\mathbb{A}_i = \{\mathbf{y}_i^c = (y_{i1}^c, \dots, y_{is_i^c}^c)^\top | y_{i1}^c \leq V_{i1}^c, \dots, y_{is_i^c}^c \leq V_{is_i^c}^c\}$ ,

$$\boldsymbol{\mu}_i = (\mathbf{X}_i^c \boldsymbol{\beta} + \mathbf{Z}_i^c \mathbf{b}_i) + \mathbf{R}_i^{co} (\mathbf{R}_i^{oo})^{-1} (\mathbf{y}_i^o - \mathbf{X}_i^o \boldsymbol{\beta} - \mathbf{Z}_i^o \mathbf{b}_i) \quad \text{e}$$

$$\mathbf{S}_i = \mathbf{R}_i^{cc} - \mathbf{R}_i^{co} (\mathbf{R}_i^{oo})^{-1} \mathbf{R}_i^{oc}.$$

Then  $\mathbf{y}_i^{(k,l)} = (y_{i1}, \dots, y_{is_i^o}, y_{is_i^o+1}^{c(k,l)}, \dots, y_{is_i}^{c(k,l)})$  is the generated sample.

**Step 2.** Sample  $\mathbf{b}_i^{(k,l)}$  from  $f(\mathbf{b}_i | \mathbf{y}_i^{(k,l)}, \kappa_i^{(k,l-1)}, \tau_i^{(k,l-1)}, \hat{\boldsymbol{\theta}}^{(k)})$ , where

$$\mathbf{b}_i | \mathbf{y}_i, \kappa_i, \tau_i \sim \text{N}_q(\boldsymbol{\Psi}_i \mathbf{Z}_i^\top \mathbf{R}_i^{-1} \kappa_i (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta}), \boldsymbol{\Psi}_i),$$

with  $\boldsymbol{\Psi} = (\kappa_i \mathbf{Z}_i^\top \mathbf{R}_i^{-1} \mathbf{Z}_i + \tau_i \mathbf{D}^{-1})^{-1}$  (Arellano-Valle et al., 2005, Lemma 2).

# SMN-MLMEC model

## SAEM - E-step

**Step 3.** Sample  $\kappa_i^{(k,l)}$  from  $f(\kappa_i | \mathbf{y}_i^{(k,l)}, \mathbf{b}_i^{(k,l)}, \tau_i^{(k,l-1)}, \widehat{\boldsymbol{\theta}}^{(k)})$ .

**Step 4.** Sample  $\tau_i^{(k,l)}$  from  $f(\tau_i | \mathbf{y}_i^{(k,l)}, \mathbf{b}_i^{(k,l)}, \kappa_i^{(k,l)}, \widehat{\boldsymbol{\theta}}^{(k)})$ .

**Observation:** Note that since  $\mathbf{y}_i | \mathbf{b}_i$  is independent of  $\tau_i$ ;  $\mathbf{b}_i$  independent of  $\kappa_i$ ; and  $\kappa_i$  and  $\tau_i$  are mutually independent, then we have

$$f(\kappa_i | \mathbf{y}_i, \mathbf{b}_i, \tau_i) \propto f(\mathbf{y}_i | \mathbf{b}_i, \kappa_i) f(\kappa_i)$$

and

$$f(\tau_i | \mathbf{y}_i, \mathbf{b}_i, \kappa_i) \propto f(\mathbf{b}_i | \tau_i) f(\tau_i).$$

# SMN-MLMEC model

## SAEM - E-step

Distribution of $\epsilon_i$	Distribution of $\kappa_i$	Distribution of $\kappa_i   \mathbf{y}_i, \mathbf{b}_i, \tau_i$
$T_{s_i}(\mathbf{0}, \mathbf{R}_i, \nu)$	Gamma( $\nu/2, \nu/2$ )	Gamma( $(\nu + s_i)/2, (D_{e_i}^2 + \nu)/2$ )
$SL_{s_i}(\mathbf{0}, \mathbf{R}_i, \nu)$	Beta( $\nu, 1$ )	TGamma( $\nu + s_i/2, D_{e_i}^2/2, 1$ )
$CN_{s_i}(\mathbf{0}, \mathbf{R}_i, \nu_1, \nu_2)$	$\nu_1 \mathbb{I}_{\{\nu_2\}}(\kappa_i) + (1 - \nu_1) \mathbb{I}_{\{1\}}(\kappa_i)$	$P(\kappa_i = \nu_2) = 1 - P(\kappa_i = 1) = p_1/p_1 + p_2$ $p_1 = \nu_1 \nu_2^{s_i/2} \exp\{-\frac{1}{2} D_{e_i}^2 \nu_2\}$ $p_2 = (1 - \nu_1) \exp\{-\frac{1}{2} D_{e_i}^2\}$
$D_{e_i}^2 = (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta} - \mathbf{Z}_i \mathbf{b}_i)^\top \mathbf{R}_i^{-1} (\mathbf{y}_i - \mathbf{X}_i \boldsymbol{\beta} - \mathbf{Z}_i \mathbf{b}_i)$		
Distribution of $\mathbf{b}_i$	Distribution of $\tau_i$	Distribution of $\tau_i   \mathbf{y}_i, \mathbf{b}_i, \kappa_i$
$T_q(\mathbf{0}, \mathbf{D}, \eta)$	Gamma( $\eta/2, \eta/2$ )	Gamma( $(\eta + q)/2, (D_{b_i}^2 + \eta)/2$ )
$SL_q(\mathbf{0}, \mathbf{D}, \eta)$	Beta( $\eta, 1$ )	TGamma( $\eta + q/2, D_{b_i}^2/2, 1$ )
$CN_q(\mathbf{0}, \mathbf{D}, \eta_1, \eta_2)$	$\eta_1 \mathbb{I}_{\{\eta_2\}}(\tau_i) + (1 - \eta_1) \mathbb{I}_{\{1\}}(\tau_i)$	$P(\tau_i = \eta_2) = 1 - P(\tau_i = 1) = q_1/q_1 + q_2$ $q_1 = \eta_1 \eta_2^{q/2} \exp\{-\frac{1}{2} D_{b_i}^2 \eta_2\}$ $q_2 = (1 - \eta_1) \exp\{-\frac{1}{2} D_{b_i}^2\}$
$D_{b_i}^2 = \mathbf{b}_i^\top \mathbf{D}^{-1} \mathbf{b}_i$		

# SMN-MLMEC model

## SAEM - E-step

► **Stochastic-approximation-step:**  $(\mathbf{y}_i^{(k,l)}, \mathbf{b}_i^{(k,l)}, \kappa_i^{(k,l)}, \tau_i^{(k,l)}), l = 1, \dots, m$ :

$$\widehat{\kappa \mathbf{y}_i^2}^{(k)} = \widehat{\kappa \mathbf{y}_i^2}^{(k-1)} + \delta_k \left( \frac{1}{m} \sum_{l=1}^m \kappa_i^{(k,l)} \mathbf{y}_i^{(k,l)} \mathbf{y}_i^{(k,l)\top} - \widehat{\kappa \mathbf{y}_i^2}^{(k-1)} \right),$$

$$\widehat{\kappa \mathbf{y}_i}^{(k)} = \widehat{\kappa \mathbf{y}_i}^{(k-1)} + \delta_k \left( \frac{1}{m} \sum_{l=1}^m \kappa_i^{(k,l)} \mathbf{y}_i^{(k,l)} - \widehat{\kappa \mathbf{y}_i}^{(k-1)} \right),$$

$$\widehat{\kappa \mathbf{b}_i^2}^{(k)} = \widehat{\kappa \mathbf{b}_i^2}^{(k-1)} + \delta_k \left( \frac{1}{m} \sum_{l=1}^m \kappa_i^{(k,l)} \mathbf{b}_i^{(k,l)} \mathbf{b}_i^{(k,l)\top} - \widehat{\kappa \mathbf{b}_i^2}^{(k-1)} \right),$$

$$\widehat{\kappa \mathbf{b}_i}^{(k)} = \widehat{\kappa \mathbf{b}_i}^{(k-1)} + \delta_k \left( \frac{1}{m} \sum_{l=1}^m \kappa_i^{(k,l)} \mathbf{b}_i^{(k,l)} - \widehat{\kappa \mathbf{b}_i}^{(k-1)} \right),$$

$$\widehat{\kappa \mathbf{y}_i \mathbf{b}_i}^{(k)} = \widehat{\kappa \mathbf{y}_i \mathbf{b}_i}^{(k-1)} + \delta_k \left( \frac{1}{m} \sum_{l=1}^m \kappa_i^{(k,l)} \mathbf{y}_i^{(k,l)} \mathbf{b}_i^{(k,l)\top} - \widehat{\kappa \mathbf{y}_i \mathbf{b}_i}^{(k-1)} \right),$$

$$\widehat{\tau \mathbf{b}_i^2}^{(k)} = \widehat{\tau \mathbf{b}_i^2}^{(k-1)} + \delta_k \left( \frac{1}{m} \sum_{l=1}^m \tau_i^{(k,l)} \mathbf{b}_i^{(k,l)} \mathbf{b}_i^{(k,l)\top} - \widehat{\tau \mathbf{b}_i^2}^{(k-1)} \right),$$

$$\widehat{\tau \mathbf{b}_i}^{(k)} = \widehat{\tau \mathbf{b}_i}^{(k-1)} + \delta_k \left( \frac{1}{m} \sum_{l=1}^m \tau_i^{(k,l)} \mathbf{b}_i^{(k,l)} - \widehat{\tau \mathbf{b}_i}^{(k-1)} \right),$$

$$\widehat{\ell h_{1i}}^{(k)} = \widehat{\ell h_{1i}}^{(k-1)} + \delta_k \left( \frac{1}{m} \sum_{l=1}^m \log h_1(\kappa_i^{(k,l)} | \hat{\nu}^{(k-1)}) - \widehat{\ell h_{1i}}^{(k-1)} \right),$$

$$\widehat{\ell h_{2i}}^{(k)} = \widehat{\ell h_{2i}}^{(k-1)} + \delta_k \left( \frac{1}{m} \sum_{l=1}^m \log h_2(\tau_i^{(k,l)} | \hat{\eta}^{(k-1)}) - \widehat{\ell h_{2i}}^{(k-1)} \right).$$

# SMN-MLMEC model

## SAEM - CM-step

Update  $\widehat{\boldsymbol{\theta}}^{(k)}$  by the maximization of  $Q(\boldsymbol{\theta}|\widehat{\boldsymbol{\theta}}^{(k)})$ , which leads to the following expressions:

$$\begin{aligned}\widehat{\boldsymbol{\beta}}^{(k+1)} &= \left( \sum_{i=1}^n \widehat{\kappa}_i^{(k)} \mathbf{x}_i^\top \widehat{\mathbf{R}}_i^{-1(k)} \mathbf{x}_i \right)^{-1} \sum_{i=1}^n \mathbf{x}_i^\top \widehat{\mathbf{R}}_i^{-1(k)} \left( \widehat{\kappa} \mathbf{y}_i^{(k)} - Z_i \widehat{\boldsymbol{\beta}}^{(k)} \right), \\ \widehat{\sigma}_{jl}^{2(k+1)} &= \begin{cases} (\sum_{i=1}^n n_i)^{-1} \sum_{i=1}^n \text{tr} \left( \widehat{\boldsymbol{\Omega}}_i^{-1(k)} \widehat{\kappa} \widehat{\boldsymbol{\epsilon}}_{ijl}^{(k)} \right) & \text{for } j = l, \\ (2 \sum_{i=1}^n n_i)^{-1} \sum_{i=1}^n \text{tr} \left[ \widehat{\boldsymbol{\Omega}}_i^{-1(k)} \left( \widehat{\kappa} \widehat{\boldsymbol{\epsilon}}_{ijl}^{(k)} + \widehat{\kappa} \widehat{\boldsymbol{\epsilon}}_{ijl}^{(k)} \right) \right] & \text{for } j \neq l, \end{cases} \\ \widehat{\boldsymbol{\phi}}^{(k+1)} &= \underset{\boldsymbol{\phi} \in (0,1) \times \mathbb{R}^+}{\text{argmax}} \left\{ -\frac{r}{2} \sum_{i=1}^n \log |\boldsymbol{\Omega}_i(\boldsymbol{\phi}, \mathbf{t}_i)| - \frac{1}{2} \sum_{i=1}^n \text{tr} \left[ \left( \widehat{\boldsymbol{\Sigma}}^{(k)} \otimes \boldsymbol{\Omega}_i(\boldsymbol{\phi}, \mathbf{t}_i) \right)^{-1} \widehat{\kappa} \mathbf{E}_i \right] \right\}, \\ \widehat{\mathbf{D}}^{(k+1)} &= \frac{1}{n} \sum_{i=1}^n \widehat{\tau} \widehat{\mathbf{b}}_i^2(k), \\ \widehat{\nu}^{(k+1)} &= \underset{\nu}{\text{argmax}} \sum_{i=1}^n \widehat{\ell} h_{1i}^{(k)}(\nu), \\ \widehat{\eta}^{(k+1)} &= \underset{\eta}{\text{argmax}} \sum_{i=1}^n \widehat{\ell} h_{2i}^{(k)}(\eta).\end{aligned}$$



# SMN-MLMEC model

The likelihood function

$$L_o(\boldsymbol{\theta}; \mathbf{y}^{obs}) = \prod_{i=1}^n \int \left[ \int_0^{\infty} f(\mathbf{y}_i | \mathbf{b}_i, \kappa_i; \boldsymbol{\theta}) h_1(\kappa_i) d\kappa_i \right] f(\mathbf{b}_i | \boldsymbol{\theta}) d\mathbf{b}_i.$$

Partitioning  $\mathbf{y}_i$ , we have

$$\begin{aligned} L_o(\boldsymbol{\theta}; \mathbf{y}^{obs}) &= \prod_{i=1}^n \int \left[ \int_0^{\infty} \phi_{s_i^o}(\mathbf{y}_i^o; \mathbf{X}_i^c \boldsymbol{\beta} - \mathbf{Z}_i^c \mathbf{b}_i, \kappa_i^{-1} \mathbf{R}_i^{oo}) \Phi_{s_i^c}(\mathbf{V}_i^c; \boldsymbol{\mu}_i, \kappa_i^{-1} \mathbf{S}_i) h_1(\kappa_i) d\kappa_i \right] \\ &\times f(\mathbf{b}_i | \boldsymbol{\theta}) d\mathbf{b}_i = \prod_{i=1}^n \int g(\mathbf{y}_i | \mathbf{b}_i, \kappa_i; \boldsymbol{\theta}) f(\mathbf{b}_i | \boldsymbol{\theta}) d\mathbf{b}_i \end{aligned} \quad (8)$$

where  $g(\mathbf{y}_i | \mathbf{b}_i, \kappa_i; \boldsymbol{\theta}) = \int_0^{\infty} \phi_{s_i^o}(\mathbf{y}_i^o; \mathbf{X}_i^c \boldsymbol{\beta} - \mathbf{Z}_i^c \mathbf{b}_i, \kappa_i^{-1} \mathbf{R}_i^{oo}) \Phi_{s_i^c}(\mathbf{V}_i^c; \boldsymbol{\mu}_i, \kappa_i^{-1} \mathbf{S}_i) h_1(\kappa_i | \nu) d\kappa_i$ . The integral involved in (8) can be compute using an importance sampling strategy. In fact, we have

$$L_o(\boldsymbol{\theta}; \mathbf{y}^{obs}) = \prod_{i=1}^n \int g(\mathbf{y}_i | \mathbf{b}_i, \kappa_i; \boldsymbol{\theta}) \frac{f(\mathbf{b}_i | \boldsymbol{\theta})}{f^*(\mathbf{b}_i | \boldsymbol{\theta})} d\mathbf{b}_i,$$

where  $f^*$  is the importance distribution. Consequently,  $L_o(\boldsymbol{\theta}; \mathbf{y}_i^{obs})$  is estimated through the following approximation

$$L_o(\boldsymbol{\theta}; \mathbf{y}^{obs}) = \prod_{i=1}^n \left[ \frac{1}{M} \sum_{m=1}^M g(\mathbf{y}_i | \mathbf{b}_{im}, \kappa_i; \boldsymbol{\theta}) \frac{f(\mathbf{b}_{im} | \boldsymbol{\theta})}{f^*(\mathbf{b}_{im} | \boldsymbol{\theta})} \right],$$

with  $\mathbf{b}_{i1}, \dots, \mathbf{b}_{im}$  being drawn from  $f^*(\mathbf{b}_i | \boldsymbol{\theta})$ .

# SMN-MLMEC model

## Model selection criteria

### ► AIC and BIC

$$\text{AIC} = 2m - 2\ell_{\max} \quad \text{and} \quad \text{BIC} = m \log N - 2\ell_{\max}.$$

### ► AIC and BIC decomposition (Zhang et al., 2014)

Let  $\mathbf{y}_{i1}^* = (\mathbf{y}_{i1}^\top, \dots, \mathbf{y}_{ir^*}^\top)^\top$  and  $\mathbf{y}_{i2}^* = (\mathbf{y}_{ir^*+1}^\top, \dots, \mathbf{y}_{ir}^\top)^\top$ , where  $\mathbf{y}_i = (\mathbf{y}_{i1}^\top, \mathbf{y}_{i2}^\top)^\top$  and  $r^* \in \{1, \dots, r\}$ , then the AIC and BIC have the following decompositions:

$$\text{AIC} = \text{AIC}_{\mathbf{y}_1^*} + \text{AIC}_{\mathbf{y}_2^* | \mathbf{y}_1^*} \quad \text{and} \quad \text{BIC} = \text{BIC}_{\mathbf{y}_1^*} + \text{BIC}_{\mathbf{y}_2^* | \mathbf{y}_1^*}.$$

### ► Assessment criteria

$$\Delta\text{AIC} = \text{AIC}_{\mathbf{y}_{2,0}^*} - \text{AIC}_{\mathbf{y}_2^* | \mathbf{y}_1^*} \quad \text{and} \quad \Delta\text{BIC} = \text{BIC}_{\mathbf{y}_{2,0}^*} - \text{BIC}_{\mathbf{y}_2^* | \mathbf{y}_1^*}.$$

# SMN-MLMEC model

A5055 clinical trial

◀ A5055

$$\begin{aligned}y_{i1k} &= \beta_{10} + \beta_{11}t_{ik} + \beta_{12}\text{treat}_i + \beta_{13}t_{ik}^{0.5} + \beta_{14}\text{treat}_i \times t_{ik} + b_{i10} + b_{i11}t_{ik} + e_{i1k}, \\y_{i2k} &= \beta_{20} + \beta_{21}t_{ik} + \beta_{22}\text{treat}_i + \beta_{23}\text{treat}_i \times t_{ik} + b_{i20} + b_{i21}t_{ik} + e_{i2k}, \\i &= 1, \dots, 44,\end{aligned}$$

- ▶  $y_{i1k}$  is the  $\log_{10}$  (RNA) outcome for subject  $i$  measured roughly at  $\text{day}_{ik}$ ;
- ▶  $y_{i2k}$  is the  $\log(\text{CD4}/\text{CD8})$  outcome for subject  $i$  measured roughly at  $\text{day}_{ik}$ ;
- ▶ 316 observations;
- ▶ 33% of all viral load measurements are below the detection limit;
- ▶  $t_{ik} = \text{day}_{ik}/7$  (week), for  $k = 1, \dots, s_i$ , where the weeks are: 0, 7, 14, 28, 56, 84, 112, 140 e 168;
- ▶  $\text{treat}_i$  is the treatment indicator (= 0 FOR treatment 1; = 1 for treatment 2);
- ▶  $b_{ij0}$  and  $b_{ij1}$  are the random intercept and random slope, respectively, for  $y_{ijk}$ ,  $j = 1, 2$ .
  
- ▶ This dataset was previously analyzed by Wang *et al.* (2015).

# SMN-MLMEC model

A5055 clinical trial

Information criteria for the *SMN-MLMEC* models under DEC structure:

	Distribution $\epsilon$ / Distribution $\mathbf{b}$								
	N/N	SL/N	T/N	N/SL	N/T	SL/SL	SL/T	T/SL	T/T
AIC	789.85	742.18	739.59	791.98	792.29	744.47	744.54	741.85	741.51
BIC	896.62	853.41	850.81	903.20	903.51	860.14	860.21	857.52	857.19

# SMN-MLMEC model

A5055 clinical trial

ML estimates with standard errors for the SMN-LMMC model under the T/N distribution:

Structure	Parameters	Estimate (SE)	Parameters	Estimate (SE)	Parameters	Estimate (SE)
DEC	$\beta_{10}$	3.743 (0.134)	$d_{11}$	0.1446 (0.0829)	$\sigma_{11}$	0.409 (0.076)
	$\beta_{11}$	0.130 (0.026)	$d_{21}$	0.0011 (0.0133)	$\sigma_{21}$	-0.039 (0.020)
	$\beta_{12}$	-0.005 (0.067)	$d_{22}$	-0.0884 (0.1182)	$\sigma_{22}$	0.050 (0.011)
	$\beta_{13}$	-0.957 (0.098)	$d_{31}$	-0.0011 (0.0033)	$\phi_1$	0.704 (0.065)
	$\beta_{14}$	-0.007 (0.025)	$d_{32}$	0.0034 (0.0027)	$\phi_2$	0.632 (0.131)
	$\beta_{20}$	-1.284 (0.077)	$d_{33}$	-0.0122 (0.0116)	$\nu$	4.737 (0.003)
	$\beta_{21}$	0.005 (0.005)	$d_{41}$	-0.0004 (0.0004)		
	$\beta_{22}$	0.252 (0.084)	$d_{42}$	0.2727 (0.0861)		
	$\beta_{23}$	-0.003 (0.007)	$d_{43}$	0.0008 (0.0015)		
				$d_{44}$	0.0001 (0.0001)	
	<i>loglik</i>	-344.79	AIC	739.59	BIC	850.81
UNC	$\beta_{10}$	3.718 (0.135)	$d_{11}$	0.4089 (0.1463)	$\sigma_{11}$	0.263 (0.053)
	$\beta_{11}$	0.129 (0.026)	$d_{21}$	-0.0112 (0.0153)	$\sigma_{21}$	-0.024 (0.012)
	$\beta_{12}$	0.003 (0.091)	$d_{22}$	-0.0964 (0.1251)	$\sigma_{22}$	0.028 (0.005)
	$\beta_{13}$	-0.955 (0.075)	$d_{31}$	0.0002 (0.0030)	$\nu$	4.340 (0.004)
	$\beta_{14}$	-0.008 (0.027)	$d_{32}$	0.0054 (0.0029)		
	$\beta_{20}$	-1.278 (0.076)	$d_{33}$	-0.0132 (0.0116)		
	$\beta_{21}$	0.005 (0.004)	$d_{41}$	-0.0006 (0.0004)		
	$\beta_{22}$	0.286 (0.081)	$d_{42}$	0.2953 (0.0785)		
	$\beta_{23}$	-0.006 (0.006)	$d_{43}$	0.0002 (0.0015)		
				$d_{44}$	0.0001 (0.0001)	
	<i>loglik</i>	-357.97	AIC	761.94	BIC	864.26

# SMN-MLMEC model

A5055 clinical trial

**Decomposition of AIC and BIC for the best *SMN-MLMEC* model:**

AIC	739.59	BIC	850.81
$AIC_{y_2^*   y_1^*}$	92.65	$BIC_{y_2^*   y_1^*}$	158.80
$AIC_{y_{2,0}^*}$	125.26	$BIC_{y_{2,0}^*}$	166.58
$\Delta AIC$	32.61	$\Delta BIC$	7.77

# Summary

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**Thank you!**